Research Article

# On the Order Statistics of Standard Normal-Based Power Method Distributions 

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Received 17 January 2012; Accepted 5 March 2012
Academic Editors: T. Y. Kam and G. Stavroulakis
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This paper derives a procedure for determining the expectations of order statistics associated with the standard normal distribution $(Z)$ and its powers of order three and five $\left(Z^{3}\right.$ and $\left.Z^{5}\right)$. The procedure is demonstrated for sample sizes of $n \leq 9$. It is shown that $Z^{3}$ and $Z^{5}$ have expectations of order statistics that are functions of the expectations for $Z$ and can be expressed in terms of explicit elementary functions for sample sizes of $n \leq 5$. For sample sizes of $n=6,7$ the expectations of the order statistics for $Z, Z^{3}$, and $Z^{5}$ only require a single remainder term.

## 1. Introduction

Order statistics have played an important role in the development of techniques associated with estimation [1, 2], hypothesis testing [3, 4], and describing data in the context of $L$ moments [5,6]. In terms of the latter, $L$-moments are based on the expectations of linear combinations of order statistics associated with a random variable X . Specifically, the first four $L$-moments are expressed as

$$
\begin{gather*}
\lambda_{1}=E\left[X_{1: 1}\right], \\
\lambda_{2}=\frac{1}{2} E\left[X_{2: 2}-X_{1: 2}\right], \\
\lambda_{3}=\frac{1}{3} E\left[X_{3: 3}-2 X_{2: 3}+X_{1: 3}\right],  \tag{1.1}\\
\lambda_{4}=\frac{1}{4} E\left[X_{4: 4}-3 X_{3: 4}+3 X_{2: 4}-X_{1: 4}\right]
\end{gather*}
$$

or more generally as

$$
\begin{equation*}
\lambda_{r}=\frac{1}{r} \sum_{j=0}^{r-1}(-1)^{j}\binom{r-1}{j} E\left[X_{r-j: r}\right] \tag{1.2}
\end{equation*}
$$

where the order statistics $X_{1: n} \leq X_{2: n} \leq \cdots \leq X_{n: n}$ are drawn from the random variable $X$. The values of $\lambda_{1}$ and $\lambda_{2}$ are measures of location and scale and are the arithmetic mean and one-half the coefficient of mean difference (or Gini's index of spread), respectively. Higherorder $L$-moments are transformed to dimensionless quantities referred to as $L$-moment ratios defined as $\tau_{r}=\lambda_{r} / \lambda_{2}$ for $r \geq 3$, and where $\tau_{3}$ and $\tau_{4}$ are the analogs to the conventional measures of skew and kurtosis. In general, $L$-moment ratios are bounded in the interval $-1<$ $\tau_{r}<1$ as is the index of $L$-skew $\left(\tau_{3}\right)$ where a symmetric distribution implies that all $L$-moment ratios with odd subscripts are zero. Other smaller boundaries can be found for more specific cases. For example, the index of $L$-kurtosis $\left(\tau_{4}\right)$ has the boundary condition for continuous distributions of [7]

$$
\begin{equation*}
\frac{5 \tau_{3}^{2}-1}{4}<\tau_{4}<1 \tag{1.3}
\end{equation*}
$$

Headrick [8] derived classes of standard normal-L-moment-based power method distributions using the polynomial transformation

$$
\begin{equation*}
p(Z)=\sum_{i=1}^{m} c_{i} Z^{i-1} \tag{1.4}
\end{equation*}
$$

where $Z \sim$ i.i.d. $N(0,1)$. Setting $m=4(m=6)$ gives the third- (fifth-) order class of power method distributions. The shape of $p(Z)$ in (1.4) is contingent on the values of the constant coefficients $c_{i}$. For the larger class of nonnormal distributions associated with $m=6$, the coefficients are computed from the system of equations given in Headrick ([8, Equations (2.8)-(2.13)] for specified values of $L$-moment ratios ( $\tau_{3, \ldots, 6}$ ). In general, $\lambda_{1}$ and $\lambda_{2}$ are standardized to the unit normal distribution as

$$
\begin{gather*}
\lambda_{1}=c_{1}+c_{3}+3 c_{5}=0 \\
\lambda_{2}=\frac{\left(4 c_{2}+10 c_{4}+43 c_{6}\right)}{4 \sqrt{\pi}}=\frac{1}{\sqrt{\pi}} . \tag{1.5}
\end{gather*}
$$

The pdf and cdf associated with (1.4) are given in parametric form as in [8, Equations (1.3) and (1.4)]

$$
\begin{align*}
f_{p(z)}(p(z)) & =\bar{f}(z)=\left(p(z), \frac{\phi(z)}{p^{\prime}(z)}\right)  \tag{1.6}\\
F_{p(z)}(p(z)) & =\bar{F}(z)=(p(z), \Phi(z))
\end{align*}
$$

where $\bar{f}: \mathfrak{R} \mapsto \mathfrak{R}^{2}$ and $\bar{F}: \mathfrak{R} \mapsto \mathfrak{R}^{2}$ are the parametric forms of the pdf and cdf with the mappings $z \mapsto(x, y)$ and $z \mapsto(x, v)$ with $x=p(z), y=\phi(z) / p^{\prime}(z), v=\Phi(z)$, and where $\phi(z)$ and $\Phi(z)$ are the standard normal pdf and cdf, respectively. For further details on the distributional properties associated with power method transformations see [9, pages 9-30] and [8] in terms of conventional moment and $L$-moment theory, respectively.

$$
\begin{gathered}
\tau_{3}=0 \\
\tau_{4}=\frac{30 \tan ^{-1}(\sqrt{2})}{\pi}-9=0.122601 \cdots
\end{gathered}
$$


(a)

$$
\begin{gathered}
\tau_{3}=0 \\
\tau_{4}=\frac{30 \tan ^{-1}(\sqrt{2})}{\pi}+\frac{\sqrt{2}}{\pi}-9=0.572759 \ldots
\end{gathered}
$$


(b)

$$
\tau_{4}=\frac{30 \tan ^{-1}(\sqrt{2})}{\pi}+\frac{385}{129 \pi \sqrt{2}}-9=0.794349 \ldots 12
$$

(c)

Figure 1: Graphs of the three standard normal-based power method distributions $p_{t}(Z)$ in (1.7) and their values of $L$-skew $\left(\tau_{3}\right)$ and $L$-kurtosis $\left(\tau_{4}\right)$.

Of concern in this study are three power method distributions related to (1.4) and (1.5) as

$$
p_{t}(Z)=c_{2 t} Z^{2 t-1}, \quad \text { where if }\left\{\begin{array}{l}
t=1,  \tag{1.7}\\
t=2, \\
t=3,
\end{array} \quad \begin{array}{l}
c_{2}=1, c_{4}=0, c_{6}=0 \\
c_{2}=0, c_{4}=2 / 5, c_{6}=0 \\
c_{2}=0, c_{4}=0, c_{6}=4 / 43
\end{array}\right.
$$

and thus $p_{1}(Z)=Z, p_{2}(Z)=(2 / 5) Z^{3}$ and $p_{3}(Z)=(4 / 43) Z^{5}$. Note that these power method distributions are symmetric and imply that $c_{1,3,5}=0$ in (1.4). The graphs of the pdfs associated with the distributions in (1.7) are given in Figure 1 along with their values of $L$ skew and $L$-kurtosis. We would point out that the importance of these distributions was noted by Stoyanov [10, page 281], ". . power transformations [such as $p_{2}(Z)$ and $p_{3}(Z)$ ] can be
considered as functional transformations on random data, usually called Box-Cox transformations. Their importance in the area of statistics and its applications is well known."

The standard normal distribution $p_{1}(Z)$ in (1.7) is the only case of the three distributions considered that is moment determinant. That is, $p_{2}(Z)$ and $p_{3}(Z)$ have the so-called classical problem of moments insofar as their respective cdfs have nonunique solutions (i.e., they are moment indeterminant, see [10-12]). However, as pointed out by Huang [12], $p_{2}(Z)$ and $p_{3}(Z)$ are determinant in the context of order statistics moments.

The derivation of the expected values of single order statistics associated with $p_{1}(Z)$ in terms of explicit elementary functions has been attempted by numerous authors (see [1317]). As indicated by Johnson et al. [18, pages 93-94] these attempts fail to give explicit expressions in terms of elementary functions for the expected values of order statistics with sample sizes of $n>5$. However, Renner [19] provides a technique for expressing the expected values of order statistics associated with $p_{1}(Z)$ for $n=6,7$ based on a single power series.

There is a paucity of research on the expectations of order statistics associated with $p_{2}(Z)$ and $p_{3}(Z)$ in the context of explicit elementary functions. Thus, what follows in Section 2 is the development of an approach for determining the expected values of the order statistics for $p_{2}(Z)$ and $p_{3}(Z)$, which is based on a generalization of Renner's [19] discussion in the context of $p_{1}(Z)$. In Section 3, some specific evaluations of the generalization are provided to demonstrate the methodology.

## 2. Methodology

The expected values of the order statistics associated with (1.7) can be determined based on the following expression [20, page 34]:

$$
\begin{align*}
& E\left[p(Z)_{j: n}\right] \\
& \quad=n 2^{-n}\binom{n-1}{j-1} \int_{0}^{\infty} p_{t}(z) \varphi(z)\left([1+\Psi(z)]^{j-1}[1-\Psi(z)]^{n-j}-[1-\Psi(z)]^{j-1}[1+\Psi(z)]^{n-j}\right) d z \tag{2.1}
\end{align*}
$$

where $p_{t}(z)$ is defined as in (1.7) and $\varphi(z)=2 \phi(z)$ and $\Psi(z)=2 \Phi(z)-1$ are the pdf and $c d f$ of the folded unit normal distribution at $z=0$. Table 1 gives a summary of some specific expansions of the polynomial in (2.1) for sample sizes of $n=1, \ldots, 9$, which are applicable to all three distributions related to $p_{t}(z)$. Inspection of Table 1 indicates that we have in general (a) $E\left[p(Z)_{j: n}\right]=-E\left[p(Z)_{n+1-j: n}\right]$, (b) the median $E\left[p(Z)_{j: n}\right]=-E\left[p(Z)_{j: n}\right]=0$, and (c) the $E\left[p(Z)_{j: n}\right]$ are linear combinations of the integrals $I_{2 r-1}$ for $r=1,2, \ldots$, with only odd subscripts appearing as only odd powers of $\Psi(z)$ appear in the polynomial expansions associated with (2.1). As such, $I_{2 r-1}$ in (2.1) can be expressed as

$$
\begin{equation*}
I_{2 r-1}=\int_{0}^{\infty} p_{t}(z) \varphi(z)[\Psi(z)]^{2 r-1} d z \tag{2.2}
\end{equation*}
$$

Equation (2.2) may be integrated by parts as

$$
\begin{equation*}
I_{2 r-1}=(2 r-1) \int_{0}^{\infty} q_{t}(z) \varphi(z)^{2}[\Psi(z)]^{2 r-2} d z \tag{2.3}
\end{equation*}
$$

Table 1: General expressions for the expected values of the order statistics for $p_{t=1,2,3}(Z)$ in (1.7) and sample sizes of $n=1, \ldots, 9$. $I_{2 r-1}$ denotes an integral in (2.1) where $r=1, \ldots, 4$.

| Sample size ( $n$ ) | Expected value |
| :---: | :---: |
| 1 | $E\left[p_{t}(Z)_{1: 1}\right]=-E\left[p_{t}(Z)_{1: 1}\right]=0$ |
| 2 | $E\left[p_{t}(Z)_{2: 2}\right]=-E\left[p_{t}(Z)_{1: 2}\right]=I_{1}=1 / \sqrt{\pi}$ |
| 3 | $\begin{aligned} & E\left[p_{t}(Z)_{2: 3}\right]=-E\left[p_{t}(Z)_{2: 3}\right]=0 \\ & E\left[p_{t}(Z)_{3: 3}\right]=-E\left[p_{t}(Z)_{1: 3}\right]=(3 / 2) I_{1}=3 /(2 \sqrt{\pi}) \end{aligned}$ |
| 4 | $\begin{aligned} & E\left[p_{t}(Z)_{3: 4}\right]=-E\left[p_{t}(Z)_{2: 4}\right]=(3 / 2)\left(I_{1}-I_{3}\right) \\ & E\left[p_{t}(Z)_{4: 4}\right]=-E\left[p_{t}(Z)_{1: 4}\right]=(1 / 2)\left(3 I_{1}+I_{3}\right) \end{aligned}$ |
| 5 | $\begin{aligned} & E\left[p_{t}(Z)_{3: 5}\right]=-E\left[p_{t}(Z)_{3: 5}\right]=0 \\ & E\left[p_{t}(Z)_{4: 5}\right]=-E\left[p_{t}(Z)_{2: 5}\right]=(5 / 2)\left(I_{1}-I_{3}\right) \\ & E\left[p_{t}(Z)_{5: 5}\right]=-E\left[p_{t}(Z)_{1: 5}\right]=(5 / 4)\left(I_{1}+I_{3}\right) \\ & \hline \end{aligned}$ |
| 6 | $\begin{aligned} & E\left[p_{t}(Z)_{4: 6}\right]=-E\left[p_{t}(Z)_{3: 6}\right]=(15 / 8)\left(I_{1}-2 I_{3}+I_{5}\right) \\ & E\left[p_{t}(Z)_{5: 6}\right]=-E\left[p_{t}(Z)_{2: 6}\right]=(15 / 16)\left(3 I_{1}-2 I_{3}-I_{5}\right) \\ & E\left[p_{t}(Z)_{6: 6}\right]=-E\left[p_{t}(Z)_{1: 6}\right]=(3 / 16)\left(5 I_{1}+10 I_{3}+I_{5}\right) \end{aligned}$ |
| 7 | $\begin{aligned} & E\left[p_{t}(Z)_{4: 7}\right]=-E\left[p_{t}(Z)_{4: 7}\right]=0 \\ & E\left[p_{t}(Z)_{5: 7}\right]=-E\left[p_{t}(Z)_{3: 7}\right]=(105 / 32)\left(I_{1}-2 I_{3}+I_{5}\right) \\ & E\left[p_{t}(Z)_{6: 7}\right]=-E\left[p_{t}(Z)_{2: 7}\right]=(21 / 8)\left(I_{1}-I_{5}\right) \\ & E\left[p_{t}(Z)_{7: 7}\right]=-E\left[p_{t}(Z)_{1: 7}\right]=(7 / 32)\left(3 I_{1}+10 I_{3}+3 I_{5}\right) \end{aligned}$ |
| 8 | $\begin{aligned} & E\left[p_{t}(Z)_{5: 8}\right]=-E\left[p_{t}(Z)_{4: 8}\right]=(35 / 16)\left(I_{1}-3 I_{3}+3 I_{5}-I_{7}\right) \\ & E\left[p_{t}(Z)_{6: 8}\right]=-E\left[p_{t}(Z)_{3: 8}\right]=(21 / 16)\left(3 I_{1}-5 I_{3}+I_{5}+I_{7}\right) \\ & E\left[p_{t}(Z)_{7: 8}\right]=-E\left[p_{t}(Z)_{2: 8}\right]=(7 / 16)\left(5 I_{1}+5 I_{3}-9 I_{5}-I_{7}\right) \\ & E\left[p_{t}(Z)_{8: 8}\right]=-E\left[p_{t}(Z)_{1: 8}\right]=(1 / 16)\left(7 I_{1}+35 I_{3}+21 I_{5}+I_{7}\right) \end{aligned}$ |
| 9 | $\begin{aligned} & E\left[p_{t}(Z)_{5: 9}\right]=-E\left[p_{t}(Z)_{5: 9}\right]=0 \\ & E\left[p_{t}(Z)_{6: 9}\right]=-E\left[p_{t}(Z)_{4: 9}\right]=(63 / 16)\left(I_{1}-3 I_{3}+3 I_{5}-I_{7}\right) \\ & E\left[p_{t}(Z)_{7: 9}\right]=-E\left[p_{t}(Z)_{3: 9}\right]=(63 / 16)\left(I_{1}-I_{3}-I_{5}+I_{7}\right) \\ & E\left[p_{t}(Z)_{8: 9}\right]=-E\left[p_{t}(Z)_{2: 9}\right]=(9 / 16)\left(3 I_{1}+7 I_{3}-7 I_{5}+3 I_{7}\right) \\ & E\left[p_{t}(Z)_{9: 9}\right]=-E\left[p_{t}(Z)_{1: 9}\right]=(9 / 32)\left(I_{1}+7 I_{3}+7 I_{5}+I_{7}\right) \end{aligned}$ |

where $q_{1}(z)=1, q_{2}(z)=(2 / 5)\left(z^{2}+2\right)$ and $q_{3}(z)=(4 / 43)\left(z^{4}+4 z^{2}+8\right)$, for $p_{1}(z), p_{2}(z)$, and $p_{3}(z)$, respectively. Note that $\Psi(0)=0$ and $\lim _{z \rightarrow+\infty} \varphi(z)=0$. Evaluating (2.3) for $r=1$ gives a coefficient of mean difference of

$$
\begin{equation*}
I_{1}=\int_{0}^{\infty} q_{t}(z) \varphi(z)^{2} d z=\frac{1}{\sqrt{\pi}} \tag{2.4}
\end{equation*}
$$

for all $p_{t}(z)$ in (1.7), which is consistent with the specification in (1.5) and given in Table 1.
The expression $[\Psi(z)]^{2 r-2}$ in (2.3) can be expressed as

$$
\begin{equation*}
[\Psi(z)]^{2 r-2}=\left(\frac{2}{\pi}\right)^{r-1}\left[\int_{0}^{z} \exp \left\{-\frac{1}{2} u^{2}\right\} d u\right]^{2 r-2} \tag{2.5}
\end{equation*}
$$

or analogously as a double integral over $\mathfrak{R}^{2}$ as

$$
\begin{equation*}
[\Psi(z)]^{2 r-2}=\left(\frac{2}{\pi}\right)^{r-1}\left[\iint_{0}^{z} \exp \left\{-\frac{1}{2}\left(z_{1}^{2}+z_{2}^{2}\right)\right\} d z_{1} d z_{2}\right]^{r-1} \tag{2.6}
\end{equation*}
$$

Using (2.6), let $z_{2}=z_{1} \tan \theta_{1}$ and thus $d z_{2}=z_{1} \sec ^{2} \theta_{1} d \theta_{1}$. Further, let $z_{1}^{2}+z_{2}^{2}=z_{1}^{2} \sec ^{2} \theta_{1}$. As such, the region of integration will be reduced to one-half of the area of the original rectangle associated with (2.6). Thus, we have

$$
\begin{align*}
{[\Psi(z)]^{2 r-2} } & =\left(\frac{2}{\pi}\right)^{r-1}\left[2 \int_{0}^{\pi / 4} \int_{0}^{z} \exp \left\{-\frac{1}{2}\left(z_{1}^{2} \sec ^{2} \theta_{1}\right)\right\} d z_{1}\left(z_{1} \sec ^{2} \theta_{1} d \theta_{1}\right)\right]^{r-1} \\
& =\left(\frac{4}{\pi}\right)^{r-1}\left[\int_{0}^{\pi / 4}\left\{\int_{0}^{z} \exp \left\{-\frac{1}{2}\left(z_{1}^{2} \sec ^{2} \theta_{1}\right)\right\} z_{1} d z_{1}\right\} \sec ^{2} \theta_{1} d \theta_{1}\right]^{r-1} \tag{2.7}
\end{align*}
$$

Subsequently, setting $z_{1}^{2}=w$ in (2.7), where $z_{1} d z_{1}=d w / 2$, gives

$$
\begin{align*}
{[\Psi(z)]^{2 r-2} } & =\left(\frac{4}{\pi}\right)^{r-1}\left[\int_{0}^{\pi / 4}\left\{\int_{0}^{z^{2}} \exp \left(-\frac{1}{2} w \sec ^{2} \theta_{1}\right) \frac{d w}{2}\right\} \sec ^{2} \theta_{1} d \theta_{1}\right]^{r-1} \\
& =\left(\frac{4}{\pi}\right)^{r-1}\left[\int_{0}^{\pi / 4}\left\{\frac{1}{2} \cdot \frac{\exp \left(-(1 / 2) w \sec ^{2} \theta_{1}\right)}{-(1 / 2) \sec ^{2} \theta_{1}}\right\}_{0}^{z^{2}} \sec ^{2} \theta_{1} d \theta_{1}\right]^{r-1} \tag{2.8}
\end{align*}
$$

and hence

$$
\begin{equation*}
[\Psi(z)]^{2 r-2}=\left(\frac{4}{\pi}\right)^{r-1}\left[\int_{0}^{\pi / 4}\left(1-\exp \left\{-\frac{1}{2}\left(z^{2} \sec ^{2} \theta_{1}\right)\right\}\right) d \theta_{1}\right]^{r-1} \tag{2.9}
\end{equation*}
$$

Expanding (2.9) yields

$$
\begin{equation*}
[\Psi(z)]^{2 r-2}=1+\left\{\sum_{k=1}^{r-1}(-1)^{k}\binom{r-1}{k}\left(\frac{4}{\pi}\right)^{k} \int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} \exp \left\{-\frac{1}{2} z^{2} \sum_{i=1}^{k} \sec ^{2} \theta_{i}\right\} d \theta_{1} \cdots d \theta_{k}\right\} \tag{2.10}
\end{equation*}
$$

where the subscript $i$ runs faster than $k$. For example, if $r=4$, then (2.10) would appear more specifically as

$$
\begin{align*}
{[\Psi(z)]^{2 r-2}=} & 1-\binom{r-1}{1}\left(\frac{4}{\pi}\right) \int_{0}^{\pi / 4} \exp \left\{-\frac{1}{2} z^{2} \sec ^{2} \theta_{1}\right\} d \theta_{1} \\
& +\binom{r-1}{2}\left(\frac{4}{\pi}\right)^{2} \iint_{0}^{\pi / 4} \exp \left\{-\frac{1}{2} z^{2}\left(\sec ^{2} \theta_{1}+\sec ^{2} \theta_{2}\right)\right\} d \theta_{1} d \theta_{2} \\
& -\binom{r-1}{3}\left(\frac{4}{\pi}\right)^{3} \iiint_{0}^{\pi / 4} \exp \left\{-\frac{1}{2} z^{2}\left(\sec ^{2} \theta_{1}+\sec ^{2} \theta_{2}+\sec ^{2} \theta_{3}\right)\right\} d \theta_{1} d \theta_{2} d \theta_{3} . \tag{2.11}
\end{align*}
$$

Substituting (2.10) into (2.3) and initially integrating with respect to $z$ (Lichtenstein, [21]) yields

$$
\begin{equation*}
\sqrt{\pi} \int_{0}^{\infty} q_{t}(z) \varphi(z)^{2} \exp \left\{-\frac{1}{2} z^{2} \sum_{i=1}^{k} \sec ^{2} \theta_{i}\right\} d z=g_{t}\left(\sec ^{2} \theta_{i}\right) \tag{2.12}
\end{equation*}
$$

where the specific forms of $g_{t}\left(\sec ^{2} \theta_{i}\right)$, which are associated with $p_{t}(z)$, are

$$
\begin{gather*}
g_{1}\left(\sec ^{2} \theta_{i}\right)=\frac{\sqrt{2}}{\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{1 / 2}} \\
g_{2}\left(\sec ^{2} \theta_{i}\right)=\frac{2 \sqrt{2}\left(5+2 \sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)}{5\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{3 / 2}}  \tag{2.13}\\
g_{3}\left(\sec ^{2} \theta_{i}\right)=\frac{4 \sqrt{2}\left(3+4\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)+8\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{2}\right)}{43\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{5 / 2}} .
\end{gather*}
$$

Equations (2.13) can be more conveniently expressed as

$$
\begin{equation*}
g_{t}\left(\sec ^{2} \theta_{i}\right)=g_{1}\left(\sec ^{2} \theta_{i}\right)-h_{t}\left(\sec ^{2} \theta_{i}\right) \tag{2.14}
\end{equation*}
$$

where the specific forms of $h_{t}\left(\sec ^{2} \theta_{i}\right)$ are

$$
\begin{gather*}
h_{1}\left(\sec ^{2} \theta_{i}\right)=0,  \tag{2.15}\\
h_{2}\left(\sec ^{2} \theta_{i}\right)=\frac{\sqrt{2}\left(\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)}{5\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{3 / 2}},  \tag{2.16}\\
h_{3}\left(\sec ^{2} \theta_{i}\right)=\frac{\sqrt{2}\left(11 \sum_{i=1}^{k} \sec ^{4} \theta_{i}+28 \sum_{i=1}^{k} \sec ^{2} \theta_{i}+22 \sum_{i<j} \sec ^{2} \theta_{i} \sec ^{2} \theta_{j}\right)}{43\left(2+\sum_{i=1}^{k} \sec ^{2} \theta_{i}\right)^{5 / 2}} \tag{2.17}
\end{gather*}
$$

and where $\sum_{i<j}$ in (2.17) indicates summing over all $k(k-1) / 2$ pairwise combinations. Hence, the integral in (2.3) can be expressed as

$$
\begin{equation*}
I_{2 r-1}=\frac{2 r-1}{\sqrt{\pi}}\left(1+\left\{\sum_{k=1}^{r-1}(-1)^{k}\binom{r-1}{k}\left(\frac{4}{\pi}\right)^{k} \int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} g_{t}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} \cdots d \theta_{k}\right\}\right) \tag{2.18}
\end{equation*}
$$

and subsequently substituting (2.14) into (2.18) gives

$$
\begin{align*}
& I_{2 r-1} \\
& =\frac{2 r-1}{\sqrt{\pi}}\left(1+\left\{\sum_{k=1}^{r-1}(-1)^{k}\binom{r-1}{k}\left(\frac{4}{\pi}\right)^{k} \int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4}\left(g_{1}\left(\sec ^{2} \theta_{i}\right)-h_{t}\left(\sec ^{2} \theta_{i}\right)\right) d \theta_{1} \cdots d \theta_{k}\right\}\right) \tag{2.19}
\end{align*}
$$

The integral associated with $g_{1}\left(\sec ^{2} \theta_{i}\right)$ in (2.19) cannot be expressed in terms of explicit elementary functions for $k>1$, which also implies $r>2$ and sample sizes of $n>5$ in Table 1. As such, we will consider the approximating function $g_{1}^{*}\left(\sec ^{2} \theta_{i}\right)$ as

$$
\begin{equation*}
g_{1}^{*}\left(\sec ^{2} \theta_{i}\right)=\left(2^{k / 2}\right) \prod_{i=1}^{k} \frac{1}{\left(2+\sec ^{2} \theta_{i}\right)^{1 / 2}} \tag{2.20}
\end{equation*}
$$

where

$$
\begin{align*}
\int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} g_{1}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} \cdots d \theta_{k} & =\int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} g_{1}^{*}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} \cdots d \theta_{k} \\
& = \begin{cases}\tan ^{-1}(1 / \sqrt{2}), & k=1 \\
0, & k \longrightarrow \infty\end{cases} \tag{2.21}
\end{align*}
$$

Thus, for finite $k>1$ we have

$$
\begin{align*}
\int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} g_{1}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} \cdots d \theta_{k} & =\int_{0}^{\pi / 4} \cdots \int_{0}^{\pi / 4} g_{1}^{*}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} \cdots d \theta_{k}+\varepsilon_{k}  \tag{2.22}\\
& =\left(\tan ^{-1}\left(\frac{1}{\sqrt{2}}\right)\right)^{k}+\varepsilon_{k}
\end{align*}
$$

where $\varepsilon_{k}$ is the remainder term required for $k>1$ and where $\varepsilon_{1}=0$ for $r=1,2$ and $n \leq 5$. Thus, using (2.22), (2.19) can be expressed as

$$
\begin{align*}
& I_{2 r-1} \\
& =\frac{2 r-1}{\sqrt{\pi}}\left(\left\{1+\sum_{k=1}^{r-1}(-1)^{k}\binom{r-1}{k}\left(\frac{4}{\pi}\right)^{k}\right.\right. \\
& \tag{2.23}
\end{align*}
$$

The remainder terms $\varepsilon_{k>1}$ in (2.23) can be solved by using (2.3), (2.15), (2.23), and the error function Erf [22], where Erf would replace $\Phi(z)$ in (2.3) where $\Psi(z)=2 \Phi(z)-1$. More specifically, Table 2 gives the values of $\varepsilon_{k}$ for $k=1, \ldots 12,25$, and 50 with 40 -digit precision.

Table 2: Computed values of the remainder term $\varepsilon_{k}$ associated with (2.23). The values were computed with 40-digit precision.

| Sample size $(n)$ | Integral | Remainder term |
| :--- | :---: | :--- |
| $1, \ldots, 5$ | $I_{1}, I_{3}$ | $\varepsilon_{1}=0.0$ |
| 6,7 | $I_{5}$ | $\varepsilon_{2}=0.03140698829552010270731937950881276500595$ |
| 8,9 | $I_{7}$ | $\varepsilon_{3}=0.05156068650031409787170392919312656858246$ |
| 10,11 | $I_{9}$ | $\varepsilon_{4}=0.05900198710355817149868423817928465212298$ |
| 12,13 | $I_{11}$ | $\varepsilon_{5}=0.05808975458203638968882522593413660371348$ |
| 14,15 | $I_{13}$ | $\varepsilon_{6}=0.05274763616761422221709626523935998463539$ |
| 16,17 | $I_{15}$ | $\varepsilon_{7}=0.04559236574104643530748593758544745949676$ |
| 18,19 | $I_{17}$ | $\varepsilon_{8}=0.03815223895234453779274127861572423887877$ |
| 20,21 | $I_{19}$ | $\varepsilon_{9}=0.03122205691467168489718556870682270636055$ |
| 22,23 | $I_{21}$ | $\varepsilon_{10}=0.02514855254614865670209122288596241803047$ |
| 24,25 | $I_{23}$ | $\varepsilon_{11}=0.02002429921405354560405588075438666460570$ |
| 26,27 | $I_{25}$ | $\varepsilon_{12}=0.01580928681263632398753707685232879723154$ |
| $\vdots$ | $\vdots$ | $\vdots$ |
| 52,53 | $I_{51}$ | $\varepsilon_{25}=0.00057455597453332805073409074487236584232$ |
| $\vdots$ | $\vdots$ | $\vdots$ |
| 102,103 | $I_{101}$ | $\varepsilon_{50}=0.00000099193614769461065745252616987082859$ |

Table 3: Expected values of order statistics for $p_{1}(Z)=Z$ for $n=4,5$.
$E\left[p_{1}(Z)_{3: 4}\right]=-\frac{3}{\sqrt{\pi}}+\frac{18 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}=0.29701138 \ldots$
$E\left[p_{1}(Z)_{4: 4}\right]=\frac{3}{\sqrt{\pi}}-\frac{6 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}=1.02937537 \ldots$
$E\left[p_{1}(Z)_{3: 5}\right]=0$
$E\left[p_{1}(Z)_{4: 5}\right]=-\frac{5}{\sqrt{\pi}}+\frac{30 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}=0.49501897 \ldots$
$E\left[p_{1}(Z)_{5: 5}\right]=\frac{5}{\sqrt{\pi}}-\frac{15 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}=1.16296447 \ldots$
Table 4: Expected values of order statistics for $p_{2}(Z)=(2 / 5) Z^{3}$ for $n=4,5$.
$E\left[p_{1}(Z)_{3: 4}\right]=-\frac{3 \sqrt{2}}{5 \pi^{3 / 2}}+E\left[p_{1}(Z)_{3: 4}\right]=0.14462665 \ldots$
$E\left[p_{2}(Z)_{4: 4}\right]=\frac{\sqrt{2}}{5 \pi^{3 / 2}}+E\left[p_{1}(Z)_{4: 4}\right]=1.08017028 \ldots$
$E\left[p_{2}(Z)_{3: 5}\right]=0$
$E\left[p_{2}(Z)_{4: 5}\right]=-\frac{\sqrt{2}}{\pi^{3 / 2}}+E\left[p_{1}(Z)_{4: 5}\right]=0.24104442 \ldots$
$E\left[p_{2}(Z)_{5: 5}\right]=\frac{1}{\sqrt{2} \pi^{3 / 2}}+E\left[p_{1}(Z)_{5: 5}\right]=1.28995174 \ldots$

Inspection of Table 2 indicates that the (positive) remainder term achieves a maximum at $\varepsilon_{4}$ and thereafter tends to zero as $k$ increases (i.e., $\varepsilon_{k} \rightarrow 0$ for $k>4$ ).

Table 5: Expected values of order statistics for $p_{3}(Z)=(4 / 43) Z^{5}$ for $n=4,5$.
$E\left[p_{3}(Z)_{3: 4}\right]=-\frac{77}{43 \sqrt{2} \pi^{3 / 2}}+E\left[p_{1}(Z)_{3: 4}\right]=0.069615569 \ldots$
$E\left[p_{3}(Z)_{4: 4}\right]=\frac{77}{129 \sqrt{2} \pi^{3 / 2}}+E\left[p_{1}(Z)_{4: 4}\right]=1.10517397 \ldots$
$E\left[p_{3}(Z)_{3: 5}\right]=0$
$E\left[p_{3}(Z)_{4: 5}\right]=-\frac{385}{129 \sqrt{2} \pi^{3 / 2}}+E\left[p_{1}(Z)_{4: 5}\right]=0.11602594 \ldots$
$E\left[p_{3}(Z)_{5: 5}\right]=\frac{385}{258 \sqrt{2} \pi^{3 / 2}}+E\left[p_{1}(Z)_{5: 5}\right]=1.35246098 \ldots$

Table 6: Expected values of order statistics for $p_{1}(Z)=Z$ for $n=6,7$.

$$
\begin{aligned}
& E\left[p_{1}(Z)_{4: 6}\right]=\frac{150 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{30 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{150 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}=0.20154683 \ldots \\
& E\left[p_{1}(Z)_{5: 6}\right]=-\frac{15}{2 \sqrt{\pi}}-\frac{75 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{60 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}-\frac{75 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}=0.64177503 \ldots \\
& E\left[p_{1}(Z)_{6: 6}\right]=\frac{15}{2 \sqrt{\pi}}+\frac{15 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{30 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{15 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}=1.26720636 \ldots \\
& E\left[p_{1}(Z)_{4: 7}\right]=0 \\
& E\left[p_{1}(Z)_{5: 7}\right]=\frac{525 \varepsilon_{2}}{2 \pi^{5 / 2}}-\frac{105 \tan ^{-1}(1 / \sqrt{2})}{2 \pi^{3 / 2}}+\frac{525 \tan ^{-1}(1 / \sqrt{2})^{2}}{2 \pi^{5 / 2}}=0.35270695 \ldots \\
& E\left[p_{1}(Z)_{6: 7}\right]=-\frac{21}{2 \sqrt{\pi}}-\frac{210 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{105 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}-\frac{210 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}=0.75737427 \ldots \\
& E\left[p_{1}(Z)_{7: 7}\right]=\frac{21}{2 \sqrt{\pi}}+\frac{105 \varepsilon_{2}}{2 \pi^{5 / 2}}-\frac{105 \tan ^{-1}(1 / \sqrt{2})}{2 \pi^{3 / 2}}+\frac{105 \tan ^{-1}(1 / \sqrt{2})^{2}}{2 \pi^{5 / 2}}=1.35217837 \ldots
\end{aligned}
$$

We would note that the approach taken here to determine $\varepsilon_{2}$ is analogous to Renner's [19] approach of developing a power series for this value. That is, the remainder term $\varepsilon_{2}$ in Table 2 is also the value approximated in [19] for $p_{1}(Z)$. Further, we would note that extending the approach in [19] for computing the remainder terms for $k>2$ would become computationally burdensome.

To demonstrate (2.23) more specifically, if $r=4$ and $t=2$ in (1.7), then the integral $I_{7}$ associated with $p_{2}(Z)$ would appear as

$$
\begin{align*}
& I_{7}=\frac{2 r-1}{\sqrt{\pi}}\left\{1-\binom{r-1}{1}\left(\frac{4}{\pi}\right)\left(\left(\tan ^{-1}\left(\frac{1}{\sqrt{2}}\right)\right)-\int_{0}^{\pi / 4} h_{2}\left(\sec ^{2} \theta_{i}\right) d \theta_{1}\right)\right. \\
&+\binom{r-1}{2}\left(\frac{4}{\pi}\right)^{2}\left(\left(\left(\tan ^{-1}\left(\frac{1}{\sqrt{2}}\right)\right)^{2}+\varepsilon_{2}\right)-\iint_{0}^{\pi / 4} h_{2}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} d \theta_{2}\right) \\
&\left.-\binom{r-1}{3}\left(\frac{4}{\pi}\right)^{3}\left(\left(\left(\tan ^{-1}\left(\frac{1}{\sqrt{2}}\right)\right)^{3}+\varepsilon_{3}\right)-\iiint_{0}^{\pi / 4} h_{2}\left(\sec ^{2} \theta_{i}\right) d \theta_{1} d \theta_{2} d \theta_{3}\right)\right\} \tag{2.24}
\end{align*}
$$

Table 7: Expected values of order statistics for $p_{2}(Z)=(2 / 5) Z^{3}$ for $n=6,7$.
$E\left[p_{2}(Z)_{4: 6}\right]=\frac{\sqrt{2}}{\pi^{3 / 2}}-\frac{10 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\pi^{5 / 2}}+E\left[p_{1}(Z)_{4: 6}\right]=0.06475951 \ldots$
$E\left[p_{2}(Z)_{5: 6}\right]=-\frac{2 \sqrt{2}}{\pi^{3 / 2}}+\frac{5 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\pi^{5 / 2}}+E\left[p_{1}(Z)_{5: 6}\right]=0.32918688 \ldots$
$E\left[p_{2}(Z)_{6: 6}\right]=\frac{2 \sqrt{2}}{\pi^{3 / 2}}-\frac{\sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\pi^{5 / 2}}+E\left[p_{1}(Z)_{6: 6}\right]=1.48210471 \ldots$
$E\left[p_{2}(Z)_{4: 7}\right]=0$
$E\left[p_{2}(Z)_{5: 7}\right]=\frac{7}{2 \sqrt{2} \pi^{3 / 2}}-\frac{35 \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{5: 7}\right]=0.11332914 \ldots$
$E\left[p_{2}(Z)_{6: 7}\right]=-\frac{7}{\sqrt{2} \pi^{3 / 2}}+\frac{14 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\pi^{5 / 2}}+E\left[p_{1}(Z)_{6: 7}\right]=0.41552998 \ldots$
$E\left[p_{2}(Z)_{7: 7}\right]=\frac{7}{2 \sqrt{2} \pi^{3 / 2}}-\frac{7 \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{\sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{7: 7}\right]=1.65986717 \ldots$

Table 8: Expected values of order statistics for $p_{3}(Z)=(4 / 43) Z^{5}$ for $n=6,7$.

$$
\begin{aligned}
& E\left[p_{3}(Z)_{4: 6}\right]=\frac{10 \sqrt{3}}{43 \pi^{5 / 2}}+\frac{385}{129 \sqrt{2} \pi^{3 / 2}}-\frac{1925 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{129 \pi^{5 / 2}}+E\left[p_{1}(Z)_{4: 6}\right]=0.02045216 \ldots \\
& E\left[p_{3}(Z)_{5: 6}\right]=-\frac{5 \sqrt{3}}{43 \pi^{5 / 2}}-\frac{385}{129 \sqrt{2} \pi^{3 / 2}}+\frac{1925 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{129 \sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{5: 6}\right]=0.16381284 \ldots \\
& E\left[p_{3}(Z)_{6: 6}\right]=\frac{\sqrt{3}}{43 \pi^{5 / 2}}+\frac{385}{129 \sqrt{2} \pi^{3 / 2}}-\frac{385 \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{129 \sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{6: 6}\right]=1.59019061 \ldots \\
& E\left[p_{3}(Z)_{4: 7}\right]=0 \\
& E\left[p_{3}(Z)_{5: 7}\right]=\frac{35 \sqrt{3}}{86 \pi^{5 / 2}}+\frac{2695}{516 \sqrt{2} \pi^{3 / 2}}-\frac{13475 \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{258 \sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{5: 7}\right]=0.03579128 \ldots \\
& E\left[p_{3}(Z)_{6: 7}\right]=-\frac{14 \sqrt{3}}{43 \pi^{5 / 2}}-\frac{2695}{258 \sqrt{2} \pi^{3 / 2}}+\frac{2695 \sqrt{2} \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{129 \pi^{5 / 2}}+E\left[p_{1}(Z)_{6: 7}\right]=0.21502146 \ldots \\
& E\left[p_{3}(Z)_{7: 7}\right]=\frac{7 \sqrt{3}}{86 \pi^{5 / 2}}+\frac{2695}{516 \sqrt{2} \pi^{3 / 2}}-\frac{2695 \tan ^{-1}(3 \sqrt{3 / 2} / 7)}{258 \sqrt{2} \pi^{5 / 2}}+E\left[p_{1}(Z)_{7: 7}\right]=1.81938546 \ldots \\
&
\end{aligned}
$$

## 3. Evaluations

Tables 3-5 give evaluations for the expected values of the order statistics for $p_{1}(Z), p_{2}(Z)$, and $p_{3}(Z)$ in (1.7), which are based on (2.23) and the general formulae given in Table 1 for sample sizes of $n=4,5$. Inspection of Tables 4 and 5 indicates that the expected values for $p_{2}(Z)$ and $p_{3}(Z)$ are all expressed in terms of elementary functions and are also functions of the expectations associated with $p_{1}(Z)$ in Table 3.

Presented in Tables 6, 7, and 8 are the evaluations for all three distributions in (1.7) for samples of sizes $n=6,7$ where the expectations of the order statistics for $p_{1}(Z), p_{2}(Z)$, and $p_{3}(Z)$ are all expressed in terms of explicit elementary functions and a single remainder term. Tables 9 and 10 give the expected values of the order statistics associated with the standard

Table 9: Expected values of order statistics for $p_{1}(Z)=Z$ for $n=8$.

$$
\begin{aligned}
E\left[p_{1}(Z)_{5: 8}\right]= & -\frac{210 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{980 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{210 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}+\frac{980 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}}=0.15251439 \ldots \\
E\left[p_{1}(Z)_{6: 8}\right]= & \frac{546 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{588 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{84 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{546 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}-\frac{588 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}}=0.47282249 \ldots \\
E\left[p_{1}(Z)_{7: 8}\right] & =-\frac{14}{\sqrt{\pi}}-\frac{462 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{196 \varepsilon_{3}}{\pi^{7 / 2}}+\frac{168 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}-\frac{462 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}+\frac{196 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}} \\
& =0.85222486 \ldots \\
E\left[p_{1}(Z)_{8: 8}\right]= & \frac{14}{\sqrt{\pi}}+\frac{126 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{28 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{84 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{126 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}-\frac{28 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}}=1.42360030 \ldots
\end{aligned}
$$

Table 10: Expected values of order statistics for $p_{1}(Z)=Z$ for $n=9$.

$$
\begin{aligned}
E\left[p_{1}(Z)_{5: 9}\right]= & 0 \\
E\left[p_{1}(Z)_{6: 9}\right]= & -\frac{378 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{1764 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{378 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}+\frac{1764 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}}=0.27452591 \ldots \\
E\left[p_{1}(Z)_{7: 9}\right]= & \frac{1008 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{1764 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{126 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{1008 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}-\frac{1764 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}}=0.57197078 \ldots \\
E\left[p_{1}(Z)_{8: 9}\right] & =-\frac{18}{\sqrt{\pi}}-\frac{882 \varepsilon_{2}}{\pi^{5 / 2}}+\frac{756 \varepsilon_{3}}{\pi^{7 / 2}}+\frac{252 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}-\frac{882 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}+\frac{756 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}} \\
& =0.93229745 \ldots \\
E\left[p_{1}(Z)_{9: 9}\right] & =\frac{18}{\sqrt{\pi}}+\frac{252 \varepsilon_{2}}{\pi^{5 / 2}}-\frac{126 \varepsilon_{3}}{\pi^{7 / 2}}-\frac{126 \tan ^{-1}(1 / \sqrt{2})}{\pi^{3 / 2}}+\frac{252 \tan ^{-1}(1 / \sqrt{2})^{2}}{\pi^{5 / 2}}-\frac{126 \tan ^{-1}(1 / \sqrt{2})^{3}}{\pi^{7 / 2}} \\
& =1.48501316 \ldots
\end{aligned}
$$

normal distribution $p_{1}(Z)$ for sample sizes of $n=8$ and $n=9$, respectively. We would also note that Mathematica [22] software is available from the authors for implementing the methodology.

## References

[1] H. L. Harter, "The use of order statistics in estimation," Operations Research, vol. 16, no. 4, pp. 783-798, 1968.
[2] N. Balakrishnan and A. C. Cohen, Order Statistics and Inference, Academic Press, Boston, Mass, USA, 1991.
[3] S. S. Shapiro and M. B. Wilk, "An analysis of variance test for normality: complete samples," Biometrika, vol. 52, pp. 591-611, 1965.
[4] Y. I. Petunin and S. A. Matveichuk, "Order statistics for testing the hypothesis of identical distribution function in two populations," Journal of Mathematical Sciences, vol. 71, no. 5, pp. 2701-2711, 1994.
[5] J. R. M. Hosking, "L-analysis and estimation of distributions using linear combinations of order statistics," Journal of the Royal Statistical Society Series B, vol. 52, no. 1, pp. 105-124, 1990.
[6] J. R. M. Hosking and J. R. Wallis, Regional Frequency Analysis: An Approach Based on L-Moments, Cambridge University Press, Cambridge, UK, 1997.
[7] M. C. Jones, "On some expressions for variance, covariance, skewness and L-moments," Journal of Statistical Planning and Inference, vol. 126, no. 1, pp. 97-106, 2004.
[8] T. C. Headrick, "A characterization of power method transformations through L-moments," Journal of Probability and Statistics, vol. 2011, Article ID 497463, 22 pages, 2011.
[9] T. C. Headrick, Statistical Simulation: Power Method Polynomials and Other Transformations, Chapman and Hall/CRC, Boca Raton, Fla, USA, 2010.
[10] J. Stoyanov, "Stieltjes classes for moment-indeterminate probability distributions," Journal of Applied Probability, vol. 41, pp. 281-294, 2004.
[11] C. Berg, "The cube of a normal distribution is indeterminate," The Annals of Probability, vol. 16, no. 2, pp. 910-913, 1988.
[12] J. S. Huang, "Moment problem of order statistics: a review," International Statistical Review, vol. 57, no. 1, pp. 59-66, 1989.
[13] H. L. Jones, "Exact lower moments of order statistics in small samples from a normal distribution," Annals of Mathematical Statistics, vol. 19, pp. 270-273, 1948.
[14] H. Ruben, "On the moments of order statistics in samples from normal populations," Biometrika, vol. 41, pp. 200-227, 1954.
[15] R. C. Bose and S. S. Gupta, "Moments of order statistics from a normal population," Biometrika, vol. 46, pp. 433-440, 1959.
[16] H. T. David, "The sample mean among the extreme normal order statistics," Annals of Mathematical Statistics, vol. 34, pp. 33-55, 1963.
[17] P. C. Joshi and N. Balakrishnan, "An identity for the moments of normal order statistics with applications," Scandinavian Actuarial Journal, no. 4, pp. 203-213, 1981.
[18] N. L. Johnson, S. Kotz, and N. Balakrishnan, Continuous Univariate Distributions, vol. 1, John Wiley and Sons, New York, NY, USA, 2nd edition, 1994.
[19] R. M. Renner, "Evaluation by power series of the means of normal order statistics of samples of sizes six and seven," Mathematical Chronicle, vol. 4, no. 2-3, pp. 141-147, 1976.
[20] H. A. David and H. N. Nagaraja, Order Statistics, John Wiley \& Sons, Hoboken, NJ, USA, 3rd edition, 2003.
[21] L. Lichtenstein, "Ueber die integration eines bestimmten integrals in bezug auf einen parameter," Gottingen Nachrichten, pp. 468-475, 1910.
[22] Wolfram Research, Inc, Mathematica, version 8.0, Wolfram Research, Inc, Champaign, Ill, USA, 2010.

